

# MA 2202 : Probability I

Satvik Saha, 19MS154, Group D

January 28, 2021

**Exercise 1** Compute the probability of obtaining a total of 10 points when two unbiased dice are rolled.

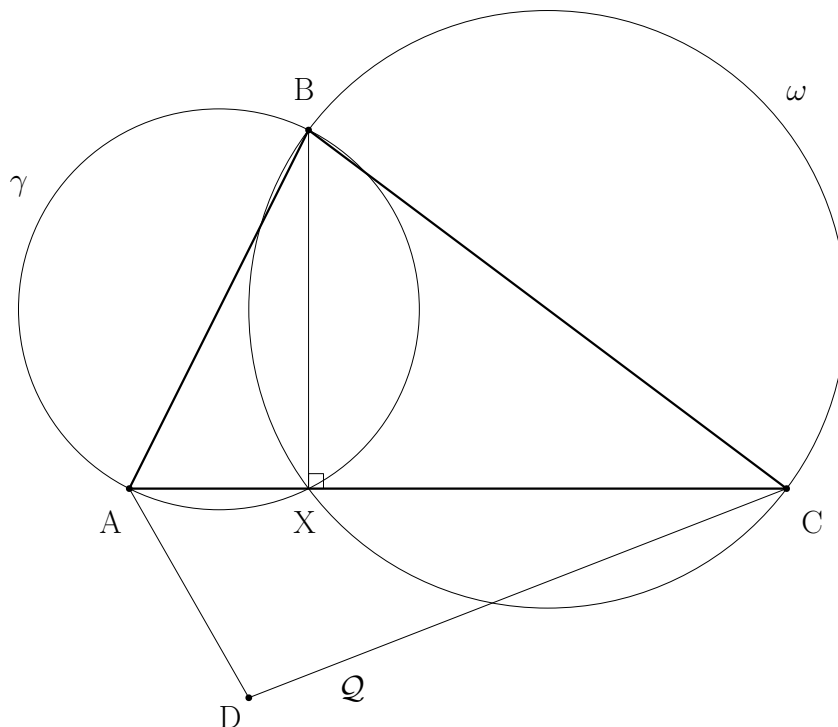
**Solution** The sample space of the possible ordered rolls is given by

$$\Omega = \{(m, n) : m, n \in \mathbb{N}, 1 \leq m, n \leq 6\}.$$

Since each of the individual rolls of a die is equally likely and the die are independent of one another, we may write that each of the outcomes  $(m, n)$  is equally likely. There are  $6 \times 6 = 36$  such tuples in  $\Omega$ , hence the probability of each of them is  $1/36$ . Note that we may set  $\mathcal{E} = \mathcal{P}(\Omega)$ . Now, we want  $m + n = 10$ , which is satisfied precisely by the tuples  $(4, 6)$   $(5, 5)$  and  $(6, 4)$ . Hence, the desired probability is simply  $3/36 = 1/12$ .

**Exercise 2** Let  $x$  be a point inside a convex quadrilateral  $\mathcal{Q}$ . Find the probability that  $x$  is neither on the boundary nor inside any of the circles drawn with the sides of the quadrilateral  $\mathcal{Q}$  as their diameters.

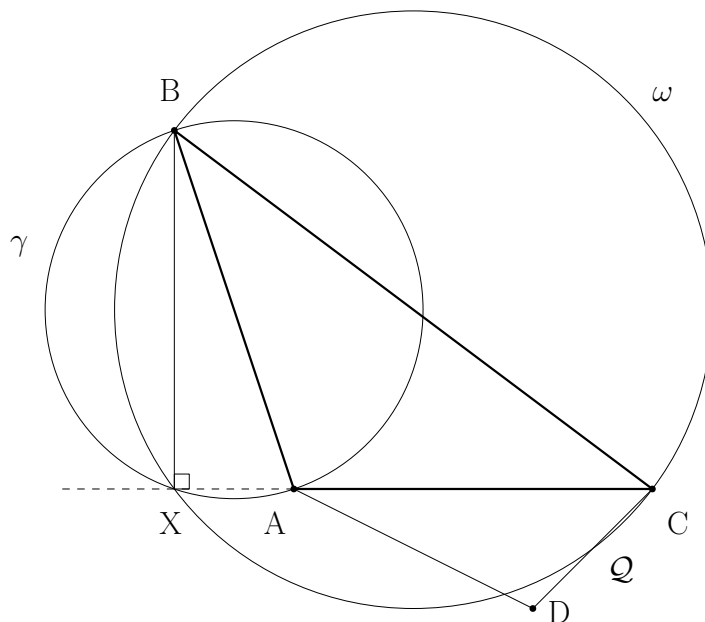
**Solution** We claim that the required probability is 0, i.e. there are no points within  $\mathcal{Q}$  which lie outside the constructed circles. To show this, pick an arbitrary point  $P \in \mathcal{Q}$ . Draw a diagonal  $AC$  of  $\mathcal{Q}$ ; the convexity of  $\mathcal{Q}$  guarantees that the segment  $AC$  lies entirely within  $\mathcal{Q}$ . Choose the triangle  $\Delta ABC$  on the side of  $AC$  where  $P$  lies, or choose the vertex  $B$  arbitrarily if  $P$  lies on the diagonal  $AC$ ; thus,  $P$  lies within (or on)  $\Delta ABC$ .



Now, construct a circle  $\gamma$  with  $AB$  as its diameter, and let it intersect the line  $AC$  at the point  $X$ . Note that the triangle  $\Delta ABX$  is inscribed within  $\gamma$ , and the side  $AB$  is a diameter of  $\gamma$ . This means that  $\Delta ABX$  is right angled, with  $\angle AXB = \pi/2$ . Thus,  $BX$  is the perpendicular through  $B$  dropped on  $AC$ . Mirroring the same argument on the other side with a circle  $\omega$  with diameter  $CB$ , we see that  $\omega$  must cut  $AC$  at the same point  $X$ , since there is only one perpendicular that can be dropped through  $B$

onto  $AC$ . Thus, the triangles  $\triangle ABX$  and  $\triangle CBX$  both lie within  $\gamma$  and  $\omega$  (including their boundaries) so the triangle  $\triangle ABC$  lies within the union of  $\gamma$  and  $\omega$ . This means that  $P$  must also lie within one of the circles. Since  $P$  and  $Q$  were arbitrary, there are no possible scenarios where  $P$  lies outside the constructed circles, hence the desired probability is 0.

Note that the given argument holds even when  $BX$  lies outside the triangle  $\triangle ABC$  – it so happens that  $\triangle ABC$  is completely within one of the triangles  $\triangle ABX$  or  $\triangle CBX$ .



It also holds for any concave quadrilateral  $Q'$ , since we can always draw a diagonal  $AC$  lying outside  $Q'$ ; now, the entire quadrilateral is contained within  $\triangle ABC$  where  $B$  is the furthest vertex from  $AC$ , and we argue as before to show that  $\triangle ABC$  lies within the circles constructed with  $AB$  and  $CB$  as diameters.

**Exercise 3** Compute the probability of getting no four consecutive heads or no four consecutive tails when a fair coin is tossed ten times.

**Solution** Let  $G_n$  denote the number of times a string of length  $n$  consisting only of the characters  $H$  and  $T$  which do not contain either of the substrings  $HHHH$  or  $TTTT$ . Note that for  $n = 1, 2, 3$ , all possible strings of which there are  $2^n$  satisfy the criterion, so  $G_1 = 2$ ,  $G_2 = 4$  and  $G_3 = 8$ .

Now, let  $G_{n,i}$  denote that number of valid strings which end in exactly  $i$  identical characters. Note that

$$G_n = G_{n,1} + G_{n,2} + G_{n,3},$$

since  $i = 4$  gives an invalid string. By adding another identical character to the end of a valid string ( $H$  if the last  $i$  characters were  $H$ ,  $T$  otherwise), we obtain  $G_{n+1,i+1} = G_{n,i}$ . By adding the complementary character to the end of a valid string ( $T$  if the last  $i$  characters were  $H$ , and vice versa), we obtain another valid string of  $n + 1$  characters with  $i = 1$ , so  $G_{n+1,1} = G_n$ . This gives a complete recursive solution :  $G_{n+3,1} = G_{n+2}$ ,  $G_{n+3,2} = G_{n+2,1} = G_{n+1}$  and  $G_{n+3,3} = G_{n+1,1} = G_n$ , so

$$G_{n+3} = G_{n+3,1} + G_{n+3,2} + G_{n+3,3} = G_{n+2} + G_{n+1} + G_n.$$

We calculate  $G_4 = 2 + 4 + 8 = 14$ ,  $G_5 = 4 + 8 + 14 = 26$ ,  $G_6 = 8 + 14 + 26 = 48$ ,  $G_7 = 14 + 26 + 48 = 88$ ,  $G_8 = 26 + 48 + 88 = 162$ ,  $G_9 = 48 + 88 + 162 = 298$ ,  $G_{10} = 88 + 162 + 298 = 548$ . Thus, there are 548 valid strings out of  $2^{10} = 1024$  strings of length 10, all of which are equally likely to occur in a coin toss experiment. This means that the desired probability is  $548/1024 = 137/256 \approx 53.5\%$ .

NOTE: The sequence  $G_n$  is simply twice the Tribonacci sequence  $0, 0, 1, 1, 2, 4, 7, 13, 24, \dots$ , with a shift. As shown here, we can give the closed form expression

$$\frac{1}{2} G_n = \frac{\alpha^{n+2}}{(\alpha - \beta)(\alpha - \gamma)} + \frac{\beta^{n+2}}{(\beta - \alpha)(\beta - \gamma)} + \frac{\gamma^{n+2}}{(\gamma - \alpha)(\gamma - \beta)},$$

where  $\alpha, \beta, \gamma$  are the roots of the polynomial  $x^3 - x^2 - x - 1$ . We supply

$$\begin{aligned}\alpha &\approx +1.8393, \\ \beta &\approx -0.4196 + 0.6063i, \\ \gamma &\approx -0.4196 - 0.6063i.\end{aligned}$$

Some code used for these calculations can be found [here](#).

The ratio  $G_{n+1}/G_n$  converges as  $n \rightarrow \infty$  to the constant  $\alpha$ . This can be seen by the fact that  $|\beta| < 1$  and  $|\gamma| < 1$ , so the growth of  $G_n$  is dominated by  $\alpha^n$ . On the other hand, the number of strings of length  $n$  grows as  $2^n$ . Thus, the probability of not getting a repeated substring of length 4 can be made as small as we want, i.e. the probability of obtaining a stretch of 4 identical tosses can be made as close to 1 by increasing  $n$ . Indeed, by  $n = 30$ , the probability of not getting a repeat is around 10%, and by  $n = 57$ , the probability is around 1%.

The corresponding limiting ratio for a repeat of  $k$  tosses always satisfies  $2(1 - 2^{-k}) < \alpha_k < 2$ , so the same argument above applies to show that we will always obtain a repeat of  $k$  tosses more often than not, for a sufficiently high number of total tosses  $n$ .

**Exercise 4** Let  $(\Omega, \mathcal{E}, P)$  be a probability space and let  $A_1, A_2, \dots, A_n \in \mathcal{E}$  with  $P(A_1 \cap \dots \cap A_n) \neq 0$ . Show that

$$P(A_1 \cap \dots \cap A_n) = P(A_1) P(A_2 | A_1) P(A_3 | A_2 \cap A_1) \cdots P(A_n | A_{n-1} \cap \dots \cap A_1).$$

**Solution** We prove the given statement by induction. The base case of  $n = 2$  demands

$$P(A_1 \cap A_2) = P(A_1) P(A_2 | A_1),$$

which is true by definition. Note that since  $P(\cap_{i=1}^n A_i) \neq 0$ , we must have  $P(\cap_{i=1}^m A_i) \neq 0$  because  $P(B) \leq P(A)$  whenever  $B \subseteq A$ .

Suppose that the statement holds for some  $n \geq 2$ . Write  $\cap_{i=1}^n A_i = \mathcal{A}$ , hence

$$P(\mathcal{A} \cap A_{n+1}) = P(\mathcal{A}) P(A_{n+1} | \mathcal{A}).$$

Expanding  $P(\mathcal{A})$  using the induction hypothesis gives the desired result.

$$P\left(\bigcap_{i=1}^{n+1} A_i\right) = P(A_1) P(A_2 | A_1) \cdots P(A_n | A_{n-1} \cap \dots \cap A_1) P(A_{n+1} | A_n \cap \dots \cap A_1).$$